#### ISSN: 2394-2975 (Online) ISSN: 2394-6814 (Print)

# An Alternative Method to Estimate The Unknown Local Parameter $\alpha$ in Laplace Distribution by Using M.L.E

# Mustfa Abduljabbar Jadah

University of Baghdad /College of Administration and Economics - Dept. of Statistics / Baghdad-IRAQ

## Abstract

W. J. Hurley (2009) using a simple induction argument that depends only on knowing the shape of a function of sums of absolute values to calculate M.L.E., An Alternative method is presented here.

# **Keywords**

MLE, median, double exponential (also referred to as Laplace distribution), the log of the random sample function that weight by The probability distribution function, likelihood function.

#### I. Introduction

W. J. Hurley (2009) in article (An Inductive Approach to calculate M.L.E.) Depends upon the function shape to determine the M.L.E, principle agree with Hurley's but it used special technique that shows more results that it couldn't be show in Hurley's; this is emphasized here

# II. Methodology

The double exponential distributionis:

$$f(x) = \frac{1}{2}e^{-|x_i - \alpha|} - \infty < x_i < \infty$$

likelihood function of double exponential distribution (Can we say the likelihood function is the log of the random sample function that weight by The double exponential function or In general it is the log of the random sample function that weight by The probability distribution function!?) is given by:

$$lnf(x_1, x_2, x_3, ..., x_n; \alpha) = -nln2 - \sum_{i=1}^{n} |x_i - \alpha|$$

$$lnf(x_1, x_2, x_3, ..., x_n; \alpha) + nln2 = -\sum_{i=1}^{n} |x_i - \alpha|$$

$$\varphi(x_1, x_2, x_3, \dots, x_n; \alpha) = -\sum_{i=1}^{n} |x_i - \alpha|$$

The function can be written as follows:

$$\varphi(x_1, x_2, x_3 \dots, x_n; \alpha) = -[\sum_{L=1}^{k} (y_L + \alpha) + \sum_{j=1}^{m} (z_j - \alpha)]$$
  
Where:

WHICE.

$$\begin{aligned} |x_i - \alpha| &= \begin{cases} (z_j - \alpha) & \text{if} x_i > \alpha \quad \text{i.e.: } z_j = x_i \\ (y_L + \alpha) & \text{if} x_i < \alpha^* \quad \text{i.e.: } y_L = -x_i \end{cases} \end{aligned}$$

\*Since parameter is not equal estimator then the equality has been deleted.

& we can write the domain of this function as:

-
$$\infty$$
< $g_1$ < $g_2$ ....< $g_m$ < $\alpha$ < $d_1$ < $d_2$ ....< $d_k$ < $\infty$  ;  $k > mork = m \ or \ k < m$  .

and

$$g_1$$
= smallest of( $-y_1, -y_2, -y_3, ..., -y_L, ..., -y_k$ ).  
 $g_2$ = second smallest of ( $-y_1, -y_2, -y_3, ..., -y_L, ..., -y_k$ ).

 $g_{\mathbf{q}} = q^{th}$  smallest of  $(-y_1, -y_2, -y_3, \dots, -y_L, \dots, -y_k)$ .

$$g_k = largest \ of (-y_1, -y_2, -y_3, ..., -y_L, ..., -y_k) \ .$$

$$d_1$$
= smallest of  $(z_1, z_2, z_3, ..., z_j, ..., z_m)$ .

 $d_2$ = second smallest of  $(z_1, z_2, z_3, ..., z_i, ..., z_m)$ .

$$d_p = p^{th} \text{ smallest of}(z_1, z_2, z_3, ..., z_j, ..., z_m)$$
.

```
\begin{array}{ll} m \text{ is } \#(\alpha > x_{\rm i}) \; ; k \text{ is } \#(\alpha < x_{\rm i}) \; ; m \; + \; k \; = \; n \; . \\ if \; \; k \; > \; m \\ \text{That is:} \\ \varphi(x_1, x_2, x_3, \ldots, x_n; \; \alpha) = -\left[\sum_{\rm L=1}^k (y_{\rm L} + \alpha) + \sum_{\rm j=1}^m (z_{\rm j} - \alpha)\right] \\ \text{we can write the domain of this function as:} \\ -\infty < g_1 < g_2, \ldots, < g_m < \alpha < d_1 < d_2, \ldots < d_k < \infty \; ; \; k \; > \; m \; . \\ \text{we note:} \\ w(m = 0, k) < w(m = 1, k) < w(m = 2, k) < \cdots < w(m = n - k, k) \; ; \; k \; > \; m \; . \\ \text{where:} \; w(m, k; \alpha) \; \text{present} \; \varphi(x_1, x_2, x_3, \ldots, x_n; \; \alpha) \; ; \; k \; > \; m \; . \\ \& \; \text{the estimator is:} \end{array}
```

if w(m = 0, k), the estimator is  $\hat{\alpha} = d_1$ .

ifw(m = 1, k), the estimator is  $\hat{\alpha} = g_1$ .

 $d_m = largest \ of(z_1, z_2, z_3, ..., z_j, ..., z_m)$ .

ifw(m = 1, k), the estimator is  $\hat{\alpha} = g_1$ . ifw(m = 2, k), the estimator is  $\hat{\alpha} = g_2$ .

$$ifw(m=n-k,k)$$
 , the estimator is  $\hat{\alpha}=g_{n-k}$  ;  $k>m$  .   
  $if$   $k=m$ 

That is:

$$\varphi(x_1, x_2, x_3, \dots, x_n; \alpha) = - [\sum_{k=1}^{k} (y_k + \alpha) + \sum_{j=1}^{m} (z_j - \alpha)]$$

& we can write the domain of this function as:

$$-\infty < g_1 < g_2 ... < g_m < \alpha < d_1 < d_2 ... < d_k < \infty; \quad k = m.$$

the estimator is 
$$\hat{\alpha} = h g_m + (1 - h)d_1$$
 for  $\alpha$ ;  $0 \le h \le 1$ ;  $k = m$ .

That is:

$$\varphi(x_1, x_2, x_3, \dots, x_n; \alpha) = - \left[ \sum_{l=1}^k (y_l + \alpha) + \sum_{j=1}^m (z_j - \alpha) \right] ;$$

& we can write the domain of this function as:

-
$$\infty$$
< $g_1$ < $g_2$ ...< $g_m$ < $\alpha$ < $d_1$ < $d_2$ ...< $d_k$ < $\infty$ ;  $k$  <  $m$  . we note:

$$w(m, k = 0) > w(m, k = 1) > w(m, k = 2) > \cdots$$
  
>  $w(m, k = n - m); k < m$ .

where:  $w(m, k; \alpha)$  present  $\varphi(x_1, x_2, x_3, ..., x_n; \alpha)$ ; k < m. & the estimator is:

ifw(m, k = 0), the estimator is  $\hat{\alpha} = g_m$ .

ifw(m, k = 1), the estimator is  $\hat{\alpha} = d_1$ if w(m, k = 2), the estimator is  $\hat{\alpha} = d_2$ 

if 
$$w(m, k = n - m)$$
, the estimator is  $\hat{\alpha} = d_{n-m}$ ;  $k < m$ .

#### ISSN: 2394-2975 (Online) ISSN: 2394-6814 (Print)

#### **III. Conclusions**

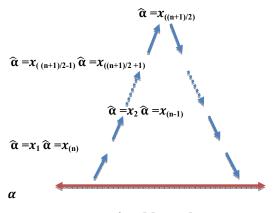
The maximum likelihood estimator for Laplace distribution to calculate  $\widehat{\alpha}$  be as follow :

$$h x_{(n/2)} + (1-h)x_{(n/2+1)} ; 0 \le h \le 1 iffn \text{ is even}$$

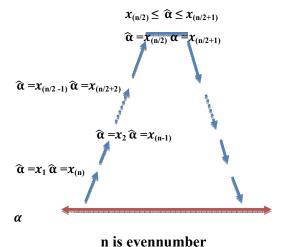
$$\therefore \widehat{\alpha} = \begin{cases} x_{((n+1)/2)} & \text{if } fn \text{ is odd} \end{cases}$$

Because the estimator make the function greatest possible &  $\{x_1, x_2, ..., x_n/m.l. \ e \ above\}$  also estimators but not make the function greatest possible &  $\widehat{\alpha} = x \ whenn = 1$ .

And the chart as: Suppose that  $x_1 < x_2 < x_3 < ... < x_{(n-2)} < x_{(n-1)} < x_{(n)}$ ;  $x_i < \alpha$  or  $x_i > \alpha$ ;  $-\infty < \alpha < \infty$ 



n is odd number



# IV. Acknowledgement

All praise for GOD highness, the compassionate, the merciful...

I would like to express my sincere gratitude to my family and my teachers and every one helped me in developing my scientific knowledge and supported me in some way or another.

My special thanks to the staff of (IJARET) (international journal of advanced research in education & technology) for their role in saving and The dissemination of science and knowledge.

### References

[1] W. J. Hurley, "An Inductive Approach to Calculate the MLE for the Double Exponential Distribution", Journal of Modern Applied Statistical Methods, 2, 594-596, 2009.

- [2] Hogg, R. V., & Craig, A. T.. "Introduction to Mathematical Statistics", (3rd Ed.), New York, NY: MacMillan Publishing Company, 1970.
- [3] Norton, R. M. The double exponential distribution: Using calculus to find a maximum likelihood n estimator, The American Statistician, 38(2), 135-136, 1984.

#### **Author's Profile**



MUSTFA ABDULJABBAR JADAH, born at April, 9th 1989, fromKut, Waset, Iraq. He has bachelor'sdegree in statistic department / College of Administration and Economics at Waset university in 2011, very good grade, and he works a day laborer in the Baghdad University/ College of Business and Economics/ Training Unit.